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Technical appendix to “Adaptive estimation of stationary Gaussian fields”

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Abstract

This is a technical appendix to “Adaptive estimation of stationary Gaussian fields” [6]. We present several proofs that have been skipped in the main paper. These proofs are organised as in Section 8 of [6].

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1. Proof of Proposition 8.1

Proof of Proposition 8.1. First, we recall the notations introduced in [3]. Let N be a positive integer. Then, \mathcal{I}_N stands for the family of subsets of $\{1, \dots, N\}$ of size less than 2. Let \mathcal{T} be a set of vectors indexed by \mathcal{I}_N . In the sequel, \mathcal{T} is assumed to be a compact subset of $\mathbb{R}^{(N(N+1)/2)+1}$. The following lemma states a slightly modified version of the upper bound in remark 7 in [3].

Lemma 1.1. *Let T be a supremum of Rademacher chaos indexed by \mathcal{I}_N of the form*

$$T := \sup_{t \in \mathcal{T}} \left| \sum_{\{i,j\}} U_i U_j t_{\{i,j\}} + \sum_{i=1}^N t_{\{i\}} + t_{\emptyset} \right|,$$

where U_1, \dots, U_N are independent Rademacher random variables. Then for any $x > 0$,

$$\mathbb{P}\{T \geq \mathbb{E}[T] + x\} \leq 4 \exp \left(-\frac{x^2}{L_1 \mathbb{E}[D]^2} \wedge \frac{x}{L_2 E} \right), \quad (1)$$

where D and E are defined by:

$$D := \sup_{t \in \mathcal{T}} \sup_{\alpha: \|\alpha\|_2 \leq 1} \left| \sum_{i=1}^N U_i \sum_{j \neq i} \alpha_j t_{\{i,j\}} \right|,$$

$$E := \sup_{t \in \mathcal{T}} \sup_{\alpha^{(1)}, \alpha^{(2)}, \|\alpha^{(1)}\|_2 \leq 1, \|\alpha^{(2)}\|_2 \leq 1} \left| \sum_{i=1}^N \sum_{j \neq i} t_{\{i,j\}} \alpha_i^{(1)} \alpha_j^{(2)} \right|.$$

Contrary to the original result of [3], the chaos are not assumed to be homogeneous. Besides, the $t_{\{i\}}$ are redundant with t_{\emptyset} . In fact, we introduced this family in order to emphasize the connection with Gaussian chaos in the next result.

A suitable application of the central limit theorem enables to obtain a corresponding bound for Gaussian chaos of order 2.

Lemma 1.2. *Let T be a supremum of Gaussian chaos of order 2.*

$$T := \sup_{t \in \mathcal{T}} \left| \sum_{\{i,j\}} t_{\{i,j\}} Y_i Y_j + \sum_i t_i Y_i^2 + t_{\emptyset} \right|, \quad (2)$$

where Y_1, \dots, Y_N are independent standard Gaussian random variable. Then, for any $x > 0$,

$$\mathbb{P}\{T \geq \mathbb{E}[T] + x\} \leq \exp\left(-\frac{x^2}{\mathbb{E}[D]^2 L_1} \wedge \frac{x}{EL_2}\right), \quad (3)$$

where

$$\begin{aligned} D &:= \sup_{t \in \mathcal{T}} \sup_{\alpha \in \mathbb{R}^N, \|\alpha\|_2 \leq 1} \sum_{i,j} Y_i (1 + \delta_{i,j}) \alpha_j t_{\{i,j\}}, \\ E &:= \sup_{t \in \mathcal{T}} \sup_{\alpha_1, \|\alpha_1\|_2 \leq 1} \sup_{\alpha_2, \|\alpha_2\|_2 \leq 1} \sum_{i,j} \alpha_{1,i} \alpha_{2,j} t_{\{i,j\}} (1 + \delta_{i,j}). \end{aligned}$$

The proof of this Lemma is postponed to the end of this section. To conclude, we derive the result of Proposition 8.1 from this last lemma. For any matrix $R \in F$, we define the vector $t^R \in \mathbb{R}^{nr(nr+1)/2+1}$ indexed by \mathcal{I}_{nr} as follows

$$t_{\{(i,k),(j,l)\}}^R := \delta_{k,l} (2 - \delta_{i,j}) \frac{R[i,j]}{n}, \quad t_{\{(i,k)\}}^R := \frac{R[i,i]}{n}, \quad \text{and } t_{\emptyset}^R := -\text{tr}(R),$$

where $\delta_{i,j}$ is the indicator function of $i = j$. In order to apply Lemma 1.2 with $N = nr$ and $\mathcal{T} = \{t^R | R \in F\}$, we have to work out the quantities D and E .

$$\begin{aligned} D &= \sup_{t^R \in \mathcal{T}} \sup_{\alpha \in \mathbb{R}^{nr}, \|\alpha\|_2 \leq 1} \left\{ \sum_{i=1}^r \sum_{k=1}^n Y[i,k] \sum_{j=1}^r \sum_{l=1}^n t_{ij}^{R,k,l} (1 + \delta_{i,j} \delta_{k,l}) \alpha_j^l \right\} \\ &= \sup_{R \in F} \sup_{\alpha \in \mathbb{R}^{nr}, \|\alpha\|_2 \leq 1} 2 \left\{ \sum_{i=1}^r \sum_{k=1}^n Y[i,k] \sum_{j=1}^r \frac{R[i,j] \alpha_j^k}{n} \right\} \\ &= \sup_{R \in F} \sup_{\alpha \in \mathbb{R}^{nr}, \|\alpha\|_2 \leq 1} \frac{2}{n} \left\{ \sum_{k=1}^n \sum_{j=1}^r \alpha_j^k \left(\sum_{i=1}^r Y[i,k] R[i,j] \right) \right\}. \end{aligned}$$

Applying Cauchy-Schwarz identity yields

$$\begin{aligned} D^2 &= \frac{4}{n^2} \sup_{R \in F} \left\{ \sum_{k=1}^n \sum_{j=1}^r \left(\sum_{i=1}^r Y_{[i,k]} R_{[i,j]} \right)^2 \right\} \\ &= \frac{4}{n} \sup_{R \in F} \text{tr}(R \overline{Y Y^*} R^*) . \end{aligned} \quad (4)$$

Let us now turn the constant E

$$\begin{aligned} E &= \sup_{t^R \in \mathcal{T}} \sup_{\substack{\alpha_1, \alpha_2 \in \mathbb{R}^{nr} \\ \|\alpha_1\|_2 \leq 1, \|\alpha_2\|_2 \leq 1}} \sum_{1 \leq i, j \leq r} \sum_{1 \leq k, l \leq n} (1 + \delta_{ij} \delta_{kl}) t_{i,j}^{R,kl} \alpha_{1,i}^k \alpha_{2,j}^l \\ &= \sup_{R \in F} \sup_{\substack{\alpha_1, \alpha_2 \in \mathbb{R}^{nr} \\ \|\alpha_1\|_2 \leq 1, \|\alpha_2\|_2 \leq 1}} \frac{2}{n} \sum_{1 \leq i, j \leq r} \sum_{1 \leq k \leq n} R_{[i,j]} \alpha_{1,i}^k \alpha_{2,j}^k . \end{aligned}$$

From this last expression, it follows that E is a supremum of L_2 operator norms

$$E = \frac{2}{n} \sup_{R \in F} \varphi_{\max} \left(\text{Diag}^{(n)}(R) \right) ,$$

where $\text{Diag}^{(n)}(R)$ is the $(nr \times nr)$ block diagonal matrix such that each diagonal block is made of the matrix R . Since the largest eigenvalue of $\text{Diag}^{(n)}(R)$ is exactly the largest eigenvalue of R , we get

$$E = \frac{2}{n} \sup_{R \in F} \varphi_{\max}(R) . \quad (5)$$

Applying Proposition 1.2 and gathering identities (4) and (5) yields

$$\mathbb{P}(Z \geq \mathbb{E}(Z) + t) \leq \exp \left[- \left(\frac{t^2}{L_1 \mathbb{E}(V)} \wedge \frac{t}{L_2 B} \right) \right] ,$$

where $B = E$ and $V = D^2$. \square

Proof of Lemma 1.1. This result is an extension of Corollary 4 in [3]. We shall closely follow the sketch of their proof adapting a few arguments. First, we upper bound the moments of $(T - \mathbb{E}(T))_+$. Then, we derive the deviation inequality from it. Here, $x_+ = \max(x, 0)$.

Lemma 1.3. *For all real numbers $q \geq 2$,*

$$\|(T - \mathbb{E}(T))_+\|_q \leq \sqrt{Lq} \mathbb{E}(D) + LqE , \quad (6)$$

where $\|T\|_q^q$ stands for the q -th moment of the random variable T . The quantities D and E are defined in Lemma 1.1.

By Lemma 1.3, for any $t \geq 0$ and any $q \geq 2$,

$$\begin{aligned} \mathbb{P}(T \geq \mathbb{E}(T) + t) &\leq \frac{\mathbb{E}[(T - \mathbb{E}(T))_+^q]}{t^q} \\ &\leq \left(\frac{\sqrt{Lq}\mathbb{E}(D) + LqE}{t} \right)^q. \end{aligned}$$

The right-hand side is at most 2^{-q} if $\sqrt{Lq}\mathbb{E}(D) \leq t/4$ and $LqE \leq t/4$. Let us set

$$q_0 := \frac{t^2}{16L\mathbb{E}(D)^2} \wedge \frac{t}{4LE}.$$

If $q_0 \geq 2$, then $\mathbb{P}(T \geq \mathbb{E}(T) + t) \leq 2^{-q_0}$. On the other hand if $q_0 < 2$, then $4 \times 2^{-q_0} \geq 1$. It follows that

$$\mathbb{P}(T \geq \mathbb{E}(T) + t) \leq 4 \exp \left(-\frac{\log(2)}{4L} \left[\frac{t^2}{4\mathbb{E}(D)^2} \wedge \frac{t}{E} \right] \right).$$

□

Proof of Lemma 1.3. This result is based on the entropy method developed in [3]. Let $f : \mathbb{R}^N \rightarrow \mathbb{R}$ be a measurable function such that $T = f(U_1, \dots, U_N)$. In the sequel, U'_1, \dots, U'_N denote independent copies of U_1, \dots, U_N . The random variable T'_i and V^+ are defined by

$$\begin{aligned} T'_i &:= f(U_1, \dots, U_{i-1}, U'_i, U_{i+1}, \dots, U_N), \\ V^+ &:= \mathbb{E} \left[\sum_{i=1}^N (T - T'_i)_+^2 | U_1^N \right], \end{aligned}$$

where U_1^N refers to the set $\{U_1, \dots, U_N\}$. Theorem 2 in [3] states that for any real $q \geq 2$,

$$\|(T - \mathbb{E}(T))_+\|_q \leq \sqrt{Lq} \|\sqrt{V^+}\|_q. \quad (7)$$

To conclude, we only have bound the moments of $\sqrt{V^+}$. By definition,

$$T = \sup_{t \in \mathcal{T}} \left| \sum_{\{i,j\}} U_i U_j t_{\{i,j\}} + \sum_{i=1}^N t_{\{i\}} + t_\emptyset \right|.$$

Since the set \mathcal{T} is compact, this supremum is achieved almost surely at an element t^0 of \mathcal{T} . For any $1 \leq i \leq N$,

$$(T - T'_i)_+^2 \leq \left((U_i - U'_i) \left| \sum_{j \neq i} U_j t^0_{\{i,j\}} \right| \right)^2.$$

Gathering this bound for any i between 1 and N , we get

$$\begin{aligned}
V^+ &\leq \sum_{i=1}^N \mathbb{E} \left[\left((U_i - U'_i) \left| \sum_{j \neq i} U_j t^0 \{i, j\} \right| \right)^2 \middle| U_1^N \right] \\
&\leq 2 \sum_{i=1}^N \left[\sum_{j \neq i} U_j t^0 \{i, j\} \right]^2 \\
&\leq 2 \sup_{\alpha \in \mathbb{R}^N, \|\alpha\|_2 \leq 1} \left[\sum_{i=1}^N \alpha_i \left(\sum_{j \neq i} t_{\{i, j\}}^0 U_j \right) \right]^2 \\
&\leq 2 \sup_{t \in \mathcal{T}} \sup_{\alpha \in \mathbb{R}^N, \|\alpha\|_2 \leq 1} \sum_{i=1}^N \left[U_i \sum_{j \neq i} \alpha_j t_{\{i, j\}} \right]^2 = 2D^2 .
\end{aligned}$$

Combining this last bound with (7) yields

$$\begin{aligned}
\|(T - \mathbb{E}(T))_+\|_q &\leq \sqrt{Lq} \sqrt{2} \|D\|_q \\
&\leq \sqrt{Lq} \left[\mathbb{E}(D) + \|(D - \mathbb{E}(D))_+\|_q \right] . \tag{8}
\end{aligned}$$

Since the random variable D defined in Lemma 1.1 is a measurable function f_2 of the variables U_1, \dots, U_N , we apply again Theorem 2 in [3].

$$\|(D - \mathbb{E}(D))_+\|_q \leq \sqrt{Lq} \left\| \sqrt{V_2^+} \right\|_q ,$$

where V_2^+ is defined by

$$V_2^+ := \mathbb{E} \left[\sum_{i=1}^N (D - D'_i)_+^2 \middle| U_i^N \right] ,$$

and $D'_i := f_2(U_1, \dots, U_{i-1}, U'_i, U_{i+1}, \dots, U_N)$. As previously, the supremum in D is achieved at some random parameter (t^0, α^0) . We therefore upper bound V_2^+ as previously.

$$\begin{aligned}
V_2^+ &\leq \sum_{i=1}^N \mathbb{E} \left[\left((U_i - U'_i) \left(\sum_{j \neq i} \alpha_j^0 t_{\{i, j\}}^0 \right) \right)^2 \middle| U_1^N \right] \\
&\leq 2 \sum_{i=1}^N \left(\sum_{j \neq i} \alpha_j^0 t_{\{i, j\}}^0 \right)^2 \\
&\leq 2 \sup_{\alpha^{(2)} \in \mathbb{R}^N, \|\alpha\|_2 \leq 1} \left(\sum_{i=1}^N \alpha_j^{(2)} \sum_{j \neq i} \alpha_i^0 t_{\{i, j\}} \right)^2 = 2E^2 .
\end{aligned}$$

Gathering this upper bound with (8) yields

$$\|(T - \mathbb{E}(T))_+\|_q \leq \sqrt{Lq} \mathbb{E}(D) + LqE .$$

□

Proof of Lemma 1.2. We shall apply the central limit theorem in order to transfer results for Rademacher chaos to Gaussian chaos. Let f be the unique function satisfying $T = f(y_1, \dots, y_N)$ for any $(y_1, \dots, y_N) \in \mathbb{R}^N$. As the set \mathcal{T} is compact, the function f is known to be continuous. Let $(U_i^{(j)})_{1 \leq i \leq N, j \geq 0}$ an i.i.d. family of Rademacher variables. For any integer $n > 0$, the random variables $Y^{(n)}$ and $T^{(n)}$ are defined by

$$\begin{aligned} Y^{(n)} &:= \left(\sum_{j=1}^n \frac{U_1^{(j)}}{\sqrt{n}}, \dots, \sum_{j=1}^n \frac{U_N^{(j)}}{\sqrt{n}} \right), \\ T^{(n)} &:= f(Y^{(n)}). \end{aligned}$$

Clearly, $T^{(n)}$ is a supremum of Rademacher chaos of order 2 with nN variables and a constant term. By the central limit theorem, $T^{(n)}$ converges in distribution towards T as n tends to infinity. Consequently, deviation inequalities for the variables $T^{(n)}$ transfer to T as long as the quantities $\mathbb{E}[D^{(n)}]$, $E^{(n)}$, and $\mathbb{E}[T^{(n)}]$ converge.

We first prove that the sequence $T^{(n)}$ converges in expectation towards T . As $T^{(n)}$ converges in distribution, it is sufficient to show that the sequence $T^{(n)}$ is asymptotically uniformly integrable. The set \mathcal{T} is compact, thus there exists a positive number t_∞ such that

$$\begin{aligned} T^{(n)} &\leq t_\infty \left[\sum_{i,j} |Y_i^{(n)} Y_j^{(n)}| + 1 \right] \\ &\leq t_\infty \left[1 + (N+1)/2 \sum_{i=1}^N \left(Y_i^{(n)} \right)^2 \right]. \end{aligned}$$

It follows that

$$\left(T^{(n)} \right)^2 \leq t_\infty^2 \left(\frac{N+1}{2} \right)^2 \frac{N+2}{2} \left[1 + \sum_{i=1}^N \left(Y_i^{(n)} \right)^4 \right]. \quad (9)$$

The sequence $Y_i^{(n)}$ does not only converge in distribution to a standard normal distribution but also in moments (see for instance [1] p.391). It follows that $\overline{\lim} \mathbb{E} \left[\left(T^{(n)} \right)^2 \right] \leq \infty$ and the sequence $f(Y^{(n)})$ is asymptotically uniformly integrable. As a consequence,

$$\lim_{n \rightarrow \infty} \mathbb{E} \left[T^{(n)} \right] = \mathbb{E}[T].$$

Let us turn to the limit of $\mathbb{E}[D^{(n)}]$. As the variable $T^{(n)}$ equals

$$T^{(n)} = \sup_{t \in \mathcal{T}} \left| \sum_{\{i,j\}} t_{\{i,j\}} \sum_{1 \leq k,l \leq n} \frac{U_i^{(k)} U_j^{(l)}}{n} + \sum_i t_i \sum_{1 \leq k \leq n} \frac{U_i^{(k)}}{\sqrt{n}} \sum_{l \neq k} \frac{U_i^{(l)}}{\sqrt{n}} + t_\emptyset + \sum_i t_i \right|,$$

it follows that

$$\begin{aligned} D^{(n)} &= \sup_{t \in \mathcal{T}} \sup_{\alpha \in \mathbb{R}^{nN}, \|\alpha\|_2 \leq 1} \left| \sum_{1 \leq i \leq N} \sum_{1 \leq k \leq n} U_i^{(k)} \left\{ \sum_{j \neq i} \frac{t_{\{i,j\}}}{n} \sum_{1 \leq l \leq n} \alpha_j^{(l)} + 2 \sum_{l \neq k} \frac{t_{\{i,l\}}}{n} \alpha_i^{(l)} \right\} \right| \\ &\leq \sup_{t \in \mathcal{T}} \sup_{\alpha \in \mathbb{R}^{nN}, \|\alpha\|_2 \leq 1} \left\{ \sum_i \frac{U_i^{(k)}}{\sqrt{n}} \sum_j (1 + \delta_{i,j}) t_{\{i,j\}} \frac{\sum_{1 \leq l \leq n} \alpha_j^{(l)}}{\sqrt{n}} \right\} + A^{(n)}, \quad (10) \end{aligned}$$

where the random variable $A^{(n)}$ is defined by

$$A^{(n)} := \sup_{t \in \mathcal{T}} \sup_{\alpha \in \mathbb{R}^{nN}, \|\alpha\|_2 \leq 1} \sum_{i=1}^N \sum_{j=1}^n t_{\{i\}} \frac{U_i^{(j)}}{n} \alpha_i^j.$$

Straightforwardly, one upper bounds $A^{(n)}$ by $t_\infty / n \sqrt{\sum_{i=1}^N \sum_{j=1}^n (U_i^{(j)})^2}$ and its expectation satisfies

$$\mathbb{E} \left(|A^{(n)}| \right) \leq t_\infty \sqrt{\frac{N}{n}},$$

which goes to 0 when n goes to infinity. Thus, we only have to upper bound the expectation of the first term in (10). Clearly, the supremum is achieved only when for all $1 \leq j \leq N$, the sequence $(\alpha_j^{(l)})_{1 \leq l \leq n}$ is constant. In such a case, the sequence $(\alpha_j^{(1)})_{1 \leq j \leq N}$ satisfies $\|\alpha^{(1)}\|_2 \leq 1/\sqrt{n}$. it follows that

$$\mathbb{E} [D^{(n)}] = \mathbb{E} \left\{ \sup_{t \in \mathcal{T}} \sup_{\alpha \in \mathbb{R}^N, \|\alpha\|_2 \leq 1} \mathbb{E} \left[\sum_i Y_i^{(n)} \sum_j (1 + \delta_{i,j}) \alpha_j \right] \right\} + \mathcal{O} \left(\frac{1}{\sqrt{n}} \right).$$

Let g be the function defined by

$$g(y_1, \dots, y_N) = \sup_{t \in \mathcal{T}} \sup_{\alpha \in \mathbb{R}^N, \|\alpha\|_2 \leq 1} \left[\sum_i y_i \sum_j (1 + \delta_{i,j}) \alpha_j \right],$$

for any $(y_1, \dots, y_N) \in \mathbb{R}^N$. The function $g(\cdot)$ is measurable and continuous as the supremum is taken over a compact set. As a consequence, $g(Y^{(n)})$ converges in distribution towards $g(Y)$. As previously, the sequence is asymptotically uniformly integrable since its moment of order 2 is uniformly upper bounded. It follows that $\lim \mathbb{E} [D^{(n)}] = \mathbb{E} [D]$.

Third, we compute the limit of $E^{(n)}$. By definition,

$$\begin{aligned} E^{(n)} &= \sup_{t \in \mathcal{T}} \sup_{\alpha_1, \alpha_2 \in \mathbb{R}^{nN}, \|\alpha_1\|_2 \leq 1, \|\alpha_2\|_2 \leq 1} \sum_{i=1}^N \sum_{k=1}^n \alpha_{1,i}^k \left[\sum_{j \neq i} \sum_{l=1}^n \alpha_{2,j}^{(l)} \frac{t_{\{i,j\}}}{n} + 2 \sum_{l \neq k} \alpha_{2,i}^{(l)} \frac{t_{\{i,l\}}}{n} \right] \\ &= \sup_{t \in \mathcal{T}} \sup_{\alpha_1, \alpha_2, \|\alpha_1\|_2 \leq 1, \|\alpha_2\|_2 \leq 1} \sum_{i=1}^N \sum_{j=1}^n (1 + \delta_{i,j}) \frac{t_{\{i,j\}}}{n} \left[\sum_{k=1}^n \sum_{l=1}^n \alpha_{1,i}^{(k)} \alpha_{2,j}^{(l)} \right] + \mathcal{O} \left(\frac{1}{n} \right). \end{aligned}$$

As for the computation of $D^{(n)}$, the supremum is achieved when the sequences $(\alpha_{1,i}^k)_{1 \leq k \leq n}$ and $(\alpha_{2,j}^l)_{1 \leq l \leq n}$ are constant for any $i \in \{1, \dots, N\}$. Thus, we only have to consider the supremum over the vectors α_1 and α_2 in \mathbb{R}^N .

$$E^{(n)} = \sup_{t \in \mathcal{T}} \sup_{\alpha_1, \alpha_2 \in \mathbb{R}^N \|\alpha_i\|_2 \leq 1} \sum_{i=1}^N \sum_{j=1}^N (1 + \delta_{ij}) t_{i,j} \alpha_{1,i} \alpha_{2,j} + \mathcal{O}\left(\frac{1}{n}\right).$$

It follows that $E^{(n)}$ converges towards E when n tends to infinity.

The random variable $T^{(n)} - \mathbb{E}(T^{(n)})$ converges in distribution towards $T - \mathbb{E}(T)$. By Lemma 1.1 ,

$$\mathbb{P}(T - \mathbb{E}(T) \geq x) \leq \liminf \exp\left(-\frac{x^2}{\mathbb{E}[D^{(n)}]^2 L_1} \wedge \frac{x}{E^{(n)} L_2}\right),$$

for any $x > 0$. Combining this upper bound with the convergence of the sequences $D^{(n)}$ and $E^{(n)}$ allows to conclude. \square

2. Proof of Theorem 3.1

Proof of Lemma 8.3. We only consider here the anisotropic case, since the isotropic case is analogous. This result is based on the deviation inequality for suprema of Gaussian chaos of order 2 stated in Proposition 8.1. For any model m' belonging to \mathcal{M} , we shall upper bound the quantities $\mathbb{E}(Z_{m'})$, $B_{m'}$, and $\mathbb{E}(W_{m'})$ defined in (42) in [6].

1. Let us first consider the expectation of $Z_{m'}$. Let $U'_{m,m'}$ be the new vector space defined by

$$U'_{m,m'} := U_{m,m'} \frac{\sqrt{D_\Sigma}}{p},$$

where $U_{m,m'}$ is introduced in the proof of Lemma 8.2 in [6]. This new space allows to handle the computation with the canonical inner product in the space of matrices. Let $\mathcal{B}_{m^2, m'^2}^{(2)}$ be the unit ball of $U'_{m,m'}$ with respect to the canonical inner product. If R belongs to $U_{m,m'}$, then $\|R\|_{\mathcal{H}'} = \|R\sqrt{D_\Sigma}/p\|_F$, where $\|\cdot\|_F$ stands for the Frobenius norm.

$$\begin{aligned} Z_{m'} &= \sup_{R \in \mathcal{B}_{m^2, m'^2}^{(2)}} \frac{1}{p^2} \text{tr} [R D_\Sigma (\overline{\mathbf{Y}\mathbf{Y}^*} - I_{p^2})] \\ &= \sup_{R \in \mathcal{B}_{m^2, m'^2}^{(2)}} \text{tr} \left[R \frac{\sqrt{D_\Sigma}}{p} (\overline{\mathbf{Y}\mathbf{Y}^*} - I_{p^2}) \right] \\ &= \left\| \Pi_{U'_{m,m'}} \frac{\sqrt{D_\Sigma}}{p} (\overline{\mathbf{Y}\mathbf{Y}^*} - I_{p^2}) \right\|_F, \end{aligned} \tag{11}$$

where $\Pi_{U'_{m,m'}}$ refers to the orthogonal projection with respect to the canonical inner product onto the space $U'_{m,m'}$. Let $F_1, \dots, F_{d_{m^2,m'/2}}$ denote an orthonormal basis of $U'_{m,m'}$.

$$\begin{aligned}
\mathbb{E}(Z_{m'}^2) &= \sum_{i=1}^{d_{m^2,m'/2}} \mathbb{E} \left[\text{tr}^2 \left(F_i \sqrt{\frac{D_\Sigma}{p^2}} (\overline{\mathbf{Y}\mathbf{Y}^*} - I_{p^2}) \right) \right] \\
&= \sum_{i=1}^{d_{m^2,m'/2}} \mathbb{E} \left[\sum_{j=1}^{p^2} F_{i[j,j]} \frac{\sqrt{D_{\Sigma[j,j]}}}{p} (\overline{\mathbf{Y}\mathbf{Y}^*}_{[j,j]} - 1) \right]^2 \\
&= \sum_{i=1}^{d_{m^2,m'/2}} \frac{2}{np^2} \text{tr}(F_i D_\Sigma F_i) \\
&\leq \sum_{i=1}^{d_{m^2,m'/2}} \frac{2\varphi_{\max}(D_\Sigma)}{np^2} = \frac{2d_{m^2,m'/2}\varphi_{\max}(\Sigma)}{np^2}.
\end{aligned}$$

Applying Cauchy-Schwarz inequality, it follows that

$$\mathbb{E}(Z_{m'}) \leq \sqrt{\frac{2d_{m^2,m'/2}\varphi_{\max}(\Sigma)}{np^2}}. \quad (12)$$

2. Using the identity (11), the quantity $B_{m'}$ equals

$$B_{m'} = \frac{2}{n} \sup_{R \in \mathcal{B}_{m^2,m'/2}^{(2)}} \varphi_{\max} \left(R \frac{\sqrt{D_\Sigma}}{p} \right).$$

As the operator norm is under-multiplicative and as it dominates the Frobenius norm, we get the following bound

$$B_{m'} \leq \frac{2\sqrt{\varphi_{\max}(\Sigma)}}{np}. \quad (13)$$

3. Let us turn to bounding the quantity $\mathbb{E}(W_{m'})$. Again, by introducing the ball $\mathcal{B}_{m^2,m'/2}^{(2)}$, we get

$$\begin{aligned}
W_{m'} &= \frac{4}{n} \sup_{R \in \mathcal{B}_{m^2,m'/2}^{(2)}} \frac{1}{p^2} \text{tr} [R \overline{\mathbf{Y}\mathbf{Y}^*} D_\Sigma R] \\
&\leq \frac{4\varphi_{\max}(\Sigma)}{np^2} \sup_{R \in \mathcal{B}_{m^2,m'/2}^{(2)}} \text{tr} [R \overline{\mathbf{Y}\mathbf{Y}^*} R] \\
&\leq \frac{4\varphi_{\max}(\Sigma)}{np^2} \left(1 + \sup_{R \in \mathcal{B}_{m^2,m'/2}^{(2)}} \text{tr} [R (\overline{\mathbf{Y}\mathbf{Y}^*} - I_{p^2}) R] \right).
\end{aligned}$$

Let $F_1, \dots, F_{d_{m^2, m'^2}}$ an orthonormal basis of $U'_{m, m'}$ and let λ be a vector in $\mathbb{R}^{d_{m^2, m'^2}}$. We write $\|\lambda\|_2$ for its L_2 norm.

$$\begin{aligned} & \mathbb{E} \left(\sup_{R \in \mathcal{B}_{m^2, m'^2}^{(2)}} \text{tr} [R (\overline{\mathbf{Y}\mathbf{Y}^*} - I_{p^2}) R]^2 \right) \\ &= \mathbb{E} \left(\sup_{\|\lambda\|_2 \leq 1} \sum_{i,j=1}^{d_{m^2, m'^2}} \lambda_i \lambda_j \text{tr} [F_i F_j (\overline{\mathbf{Y}\mathbf{Y}^*} / n - I_{p^2})] \right)^2 \\ &\leq \sum_{i,j=1}^{d_{m^2, m'^2}} \mathbb{E} \left(\text{tr} [F_i F_j (\mathbf{Y}\mathbf{Y}^* / n - I_{p^2})]^2 \right). \end{aligned}$$

The second inequality is a consequence of Cauchy-Schwarz inequality in $\mathbb{R}^{(d_{m^2, m'^2})^2}$ since the l_2 norm of the vector $(\lambda_i \lambda_j)_{1 \leq i, j \leq d_{m^2, m'^2}} \in \mathbb{R}^{d_{m^2, m'^2}^2}$ is bounded by 1. Since the matrices F_i are diagonal, we get

$$\mathbb{E} \left(\sup_{R \in \mathcal{B}_{m^2, m'^2}^{(2)}} \text{tr} [R (\mathbf{Y}\mathbf{Y}^* / n - I) R]^2 \right) \leq \frac{2}{n} \sum_{i,j=1}^{d_{m^2, m'^2}} \|F_i F_j\|_2^2.$$

It remains to bound the norm of the products $F_i F_j$ for any i, j between 1 and d_{m^2, m'^2} .

$$\sum_{i,j=1}^{d_{m^2, m'^2}} \|F_i F_j\|_2^2 = \sum_{i,j=1}^{d_{m^2, m'^2}} \sum_{k=1}^{p^2} F_i[k, k]^2 F_j[k, k]^2 = \sum_{k=1}^{p^2} \left(\sum_{i=1}^{d_{m^2, m'^2}} F_i[k, k]^2 \right)^2.$$

For any $k \in \{1, \dots, p^2\}$, $\sum_{i=1}^{d_{m^2, m'^2}} F_i[k, k]^2 \leq 1$ since $(F_1, \dots, F_{d_{m^2, m'^2}})$ form an orthonormal family. Hence, we get

$$\sum_{i,j=1}^{d_{m^2, m'^2}} \|F_i F_j\|_2^2 \leq \sum_{k=1}^{p^2} \sum_{i=1}^{d_{m^2, m'^2}} F_i[k, k]^2 = d_{m^2, m'^2}.$$

All in all, we have proved that

$$\mathbb{E}(W_{m'}) \leq \frac{4\varphi_{\max}(\Sigma)}{np^2} \left[1 + \sqrt{\frac{2d_{m^2, m'^2}}{n}} \right]. \quad (14)$$

Gathering these three bounds and applying Proposition 8.1 allows to obtain the following deviation inequality:

$$\begin{aligned} & \mathbb{P} \left(Z_{m'} \geq \sqrt{\frac{2\varphi_{\max}(\Sigma)}{n}} \left\{ \sqrt{1 + \alpha/2} \sqrt{d_{m^2, m'^2}} + \xi \right\} \right) \\ &\leq \exp \left\{ - \left[\frac{[(\sqrt{1+\alpha/2}-1)\sqrt{d_{m^2, m'^2}}+\xi]^2}{2L_1(1+\sqrt{2d_{m^2, m'^2}/n})} \wedge \frac{\sqrt{n}[(\sqrt{1+\alpha/2}-1)\sqrt{d_{m^2, m'^2}}+\xi]}{\sqrt{2}L_2} \right] \right\} \\ &\leq \exp \left\{ - \left[\frac{\omega_{m, m'}^2}{2L_1(1+\sqrt{2d_{m^2, m'^2}/n})} \wedge \frac{\sqrt{n}\omega_{m, m'}}{\sqrt{2}L_2} \right] - \left[\frac{\xi\omega_{m, m'}}{L_1[1+\sqrt{2d_{m^2, m'^2}/n}]} \wedge \frac{\sqrt{n}\xi}{\sqrt{2}L_2} \right] \right\}, \end{aligned}$$

where $\omega_{m,m'} = \left(\sqrt{1+\alpha/2}-1\right) \sqrt{d_{m^2,m'^2}}$. As n and d_{m^2,m'^2} are larger than one, there exists a universal constant L'_2 such that

$$\left[\frac{(\sqrt{1+\alpha/2}-1)^2 d_{m^2,m'^2}}{2L_1(1+\sqrt{2d_{m^2,m'^2}/n})} \wedge \frac{\sqrt{n}(\sqrt{1+\alpha/2}-1)\sqrt{d_{m^2,m'^2}}}{\sqrt{2}L_2} \right] \geq 4L'_2 \sqrt{d_{m^2,m'^2}} \left[\left(\sqrt{1+\alpha/2}-1\right)^2 \wedge \left(\sqrt{1+\alpha/2}-1\right) \right].$$

Since the vector space $U_{m,m'}$ contains all the matrices $D(\theta')$ with θ' belonging to m' , d_{m^2,m'^2} is larger than $d_{m'}$. Besides, by concavity of the square root function, it holds that $\sqrt{1+\alpha/2}-1 \geq \alpha[4\sqrt{1+\alpha/2}]^{-1}$. Setting $L'_1 := [4L_1(1+\sqrt{2})]^{-1} \wedge [\sqrt{2}L_2]^{-1}$ and arguing as previously leads to

$$\frac{\xi(\sqrt{1+\alpha/2}-1)\sqrt{d_{m^2,m'^2}}}{L_1(1+\sqrt{2d_{m^2,m'^2}/n})} \wedge \frac{\sqrt{n}\xi}{\sqrt{2}L_2} \geq L'_1 \xi \left[\frac{\alpha}{\sqrt{1+\alpha/2}} \wedge \sqrt{n} \right].$$

Gathering these two inequalities allows us to conclude that

$$\begin{aligned} & \mathbb{P} \left(Z_{m'} \geq \sqrt{\frac{2\varphi_{\max}(\Sigma)}{n}} \left\{ \sqrt{(1+\alpha/2)d_{m^2,m'^2}} + \xi \right\} \right) \\ & \leq \exp \left\{ -L'_2 \sqrt{d_{m'}} \left(\frac{\alpha}{\sqrt{1+\alpha/2}} \wedge \frac{\alpha^2}{1+\alpha/2} \right) - L'_1 \xi \left[\frac{\alpha}{\sqrt{1+\alpha/2}} \wedge \sqrt{n} \right] \right\}. \end{aligned}$$

□

Proof of Lemma 8.4 in [6]. The approach falls in two parts. First, we relate the dimensions d_m and d_{m^2} to the number of nodes of the torus Λ that are closer than r_m or $2r_m$ to the origin $(0,0)$. We recall that the quantity r_m is introduced in Definition 2.1 of [6]. Second, we compute a nonasymptotic upper bound of the number of points in \mathbb{Z}^2 that lie in the disc of radius r . This second step is quite tedious and will only give the main arguments.

Let m be a model of the collection \mathcal{M}_1 . By definition, m is the set of points lying in the disc of radius r_m centered on $(0,0)$. Hence,

$$\Theta_m = \text{vect} \{ \Psi_{i,j}, (i,j) \in m \},$$

where the matrices $\Psi_{i,j}$ are defined by Eq. (14) in [6]. As $\Psi_{i,j} = \Psi_{-i,-j}$, the dimension d_m of Θ_m is exactly the number of orbits of m under the action of the central symmetry s .

As d_{m^2} is defined as the dimension of the space U_m , it also corresponds to the dimension of the space

$$\text{vect} \{ C(\theta), \theta \in \Theta_m \} + \text{vect} \{ C(\theta)^2, \theta \in \Theta_m \}, \quad (15)$$

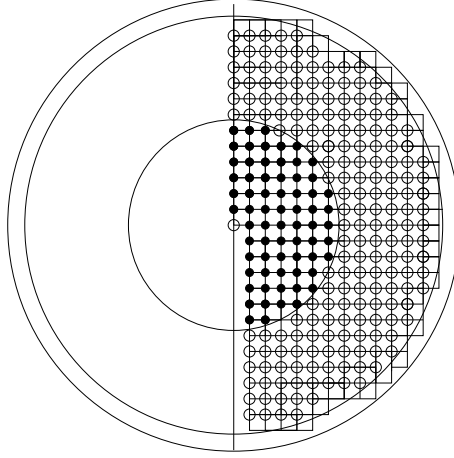


FIGURE 1. The black dots represent the orbit space of m and the white dots represent the remaining points of the orbit space of $\mathcal{N}(m)$.

which is clearly in one to one correspondence with U_m . Straightforward computations lead to the following identity:

$$\begin{aligned} C(\Psi_{i_1, j_1})C(\Psi_{i_2, j_2}) &= C(\Psi_{i_1+i_2, j_1+j_2}) [1 + s_{i_1+i_2, j_1+j_2}] \\ &+ C(\Psi_{i_1-i_2, j_1-j_2}) [1 + s_{i_1-i_2, j_1-j_2}] , \end{aligned}$$

where $s_{x,y}$ is the indicator function of $x = -x$ and $y = -y$ in the torus Λ . Combining this property with the definition of Θ_m , we embed the space (15) in the space

$$\text{vect} \{ C(\Psi_{i_1+i_2, j_1+j_2}), (i_1, j_1), (i_2, j_2) \in m \cup \{(0, 0)\} \} ,$$

and this last space is in one to one correspondence with

$$\text{vect} \{ \Psi_{i_1+i_2, j_1+j_2}, (i_1, j_1), (i_2, j_2) \in m \cup \{(0, 0)\} \} . \quad (16)$$

In the sequel, $\mathcal{N}(m)$ stands for the set

$$\{(i_1 + i_2, j_1 + j_2), (i_1, j_1), (i_2, j_2) \in m \cup \{(0, 0)\} \} .$$

Thus, the dimension d_{m^2} is smaller or equal to the number of orbits of $\mathcal{N}(m)$ under the action of the symmetry s .

To conclude, we have to compare the number of orbits in m and the number of orbits in $\mathcal{N}(m)$. We distinguish two cases depending whether $2r_m + 1 \leq p$ or $2r_m + 1 > p$. First, we assume that $2r_m + 1 \leq p$. For such values the disc of radius r_m centered on the points $(0, 0)$ in not overlapping itself on the torus except on a set of null Lebesgue measure. In the sequel, $\lfloor x \rfloor$ refers to the largest integer smaller than x . We represent the orbit space of m as in Figure 1. To any of these points, we associate a square of size 1. If we add $2 + 2\lfloor r_m \rfloor$ squares to

the d_m first squares, we remark that the half disc centered on $(0,0)$ and with length r_m is contained in the reunion of these squares. Then, we get

$$d_m + 2 + 2\lfloor r_m \rfloor \geq \frac{\pi r_m^2}{2} . \quad (17)$$

The points in $\mathcal{N}(m)$ are closer than $2r_m$ from the origin. Consequently, all the squares associated to representants of $\mathcal{N}(m)$ are included in the disc of radius $2r_m + \sqrt{2}$.

$$d_{m^2} + 2 + 2\lfloor 2r_m \rfloor \leq \frac{\pi}{2} \left\{ 2r_m + \sqrt{2} \right\}^2 .$$

Combining these two inequalities, we are able to upper bound d_{m^2}

$$\begin{aligned} 2 + 2\lfloor 2r_m \rfloor + d_{m^2} &\leq 4 \left\{ 1 + \frac{\sqrt{2}}{2r_m} \right\}^2 (d_m + 1 + 2\lfloor r_m \rfloor) , \\ d_{m^2} &\leq 4 \left\{ 1 + \frac{\sqrt{2}}{2r_m} \right\}^2 d_m + 4 \left\{ 1 + \frac{\sqrt{2}}{2r_m} \right\}^2 (1 + 2\lfloor r_m \rfloor) . \end{aligned}$$

Applying again inequality (17), we upper bound r_m :

$$r_m \leq \frac{2}{\pi} \left[1 + \sqrt{1 + \frac{\pi}{2}(1 + d_m)} \right] .$$

Gathering these two last bounds yields

$$d_{m^2} \leq 4 \left\{ 1 + \frac{\sqrt{2}}{2r_m} \right\}^2 \left[1 + \frac{1}{d_m} \left(1 + \frac{4}{\pi} \left[1 + \sqrt{1 + \frac{\pi}{2}(1 + d_m)} \right] \right) \right] d_m .$$

This upper bound is equivalent to $4d_m$, when d_m goes to infinity. Computing the ratio d_{m^2}/d_m for every model m of small dimension allows to conclude.

Let us turn to the case $2r_m + 1 > p$. Suppose that p is larger or equal to 9. The lower bound (17) does not necessarily hold anymore. Indeed, the disc is overlapping with itself because of toroidal effects. Nevertheless, we obtain a similar lower bound by replacing r_m by $(p-1)/2$:

$$d_m + 2 + 2\lfloor \frac{p-1}{2} \rfloor \geq \frac{\pi(p-1)^2}{8} .$$

The number of orbits of Λ under the action of the symmetry s is $(p^2 + 1)/2$ if p is odd and $[(p+1)^2 - 1]/2$ if p is even. It follows that $d_{m^2} \leq [(p+1)^2 - 1]/2$. Gathering these two bounds, we get

$$\frac{d_{m^2}}{d_m} \leq \frac{(p+1)^2}{\pi(p-1)^2/4 - 2(p+1)} .$$

This last quantity is smaller than 4 for any $p \geq 9$. An exhaustive computation of the ratios when $p < 9$ allows to conclude.

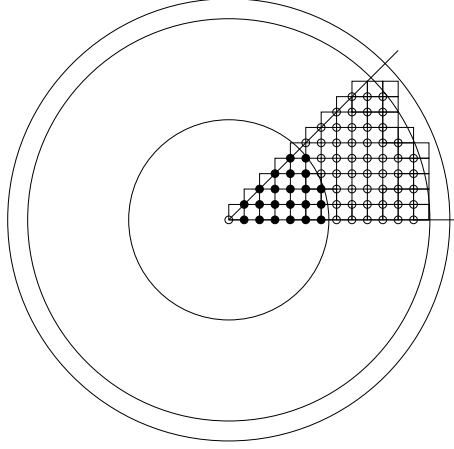


FIGURE 2. The black dots represent the orbit space of m under the action of G and the white dots represent the remaining points of the orbit space of $\mathcal{N}^{\text{iso}}(m)$.

Let us turn to the isotropic case. Arguing as previously, we observe that the dimension d_m^{iso} is the number of orbits of the set m under the action of the group G introduced in in [6] Sect.1.1 whereas d_{m^2} is smaller or equal to the number of orbits of $\mathcal{N}^{\text{iso}}(m)$ under the action of G . As for anisotropic models, we choose represent these orbits on the torus and associate squares of size 1 (see Figure 2). Assuming that $r_m < (p-1)/2$, we bound d_m and d_{m^2} .

$$\begin{aligned} d_m + 1 &\geq \frac{1}{8}\pi r_m^2 + \frac{1}{2}\left\lfloor \frac{\sqrt{2}r_m}{2} \right\rfloor, \\ d_{m^2} &\leq 4 \left\{ 1 + \frac{\sqrt{2}}{2r_m} \right\}^2 \frac{1}{8}\pi r_m^2 + \frac{1}{2}\left\lfloor \sqrt{2}r_m \right\rfloor. \end{aligned}$$

Gathering these two inequalities, we get

$$d_{m^2} \leq 4 \left\{ 1 + \frac{\sqrt{2}}{2r_m} \right\}^2 d_m.$$

As a consequence, d_{m^2} is smaller than $4d_m$ when d_m goes to infinity. As previously, computing the ratio d_{m^2}/d_m for models m of small dimension allows to conclude. The case $r_m > (p-1)/2$ is handled as for the anisotropic case. \square

3. Proofs of the minimax bounds

Proof of Lemma 8.5 in [6]. This lower bound is based on an application of Fano's approach. See [7] for a review of this method and comparisons with Le Cam's and

Assouad's Lemma. The proof follows three main steps: First, we upper bound the Kullback-Leibler entropy between distributions corresponding to θ_1 and θ_2 in the hypercube. Second, we find a set of points in the hypercube well separated with respect to the Hamming distance. Finally, we conclude by applying Birgé's version of Fano's lemma.

Lemma 3.1. *The Kullback-Leibler entropy between two mean zero-Gaussian vectors of size p^2 with precision matrices $(I_{p^2} - C(\theta_1)) / \sigma^2$ and $(I_{p^2} - C(\theta_2)) / \sigma^2$ equals*

$$\mathcal{K}(\theta_1, \theta_2) = 1/2 \left[\log \left(\frac{|I_{p^2} - C(\theta_1)|}{|I_{p^2} - C(\theta_2)|} \right) + \text{tr} \left([I_{p^2} - C(\theta_2)] [I_{p^2} - C(\theta_1)]^{-1} \right) - p^2 \right],$$

where for any square matrix A , $|A|$ refers to the determinant of A .

This statement is classical and its proof is omitted. The matrices $(I_{p^2} - C(\theta_1))$ and $(I_{p^2} - C(\theta_2))$ are diagonalizable in the same basis since they are symmetric block circulant (Lemma A.1 in [6]). Transforming vectors of size p^2 into $p \times p$ matrices, we respectively define λ_1 and λ_2 as the $p \times p$ matrices of eigenvalues of $(I_{p^2} - C(\theta_1))$ and $(I_{p^2} - C(\theta_2))$. It follows that

$$\mathcal{K}(\theta_1, \theta_2) = 1/2 \sum_{1 \leq i, j \leq p} \left(\frac{\lambda_{2[i,j]}}{\lambda_{1[i,j]}} - \log \left(\frac{\lambda_{2[i,j]}}{\lambda_{1[i,j]}} \right) - 1 \right).$$

For any $x > 0$, the following inequality holds

$$x - 1 - \log(x) \leq \frac{9}{64} \left(x - \frac{1}{x} \right)^2.$$

It is easy to establish by studying the derivative of corresponding functions. As a consequence,

$$\begin{aligned} \frac{\lambda_{2[i,j]}}{\lambda_{1[i,j]}} - \log \left(\frac{\lambda_{2[i,j]}}{\lambda_{1[i,j]}} \right) - 1 &\leq \frac{9}{64} \left(\frac{\lambda_{2[i,j]}}{\lambda_{1[i,j]}} - \frac{\lambda_{1[i,j]}}{\lambda_{2[i,j]}} \right)^2 \\ &\leq \frac{9}{64} \left(\frac{1}{\lambda_{1[i,j]}} + \frac{1}{\lambda_{2[i,j]}} \right)^2 (\lambda_{1[i,j]} - \lambda_{2[i,j]})^2. \end{aligned} \quad (18)$$

Let us first consider the anisotropic case. Let m be a model in \mathcal{M}_1 and let θ' belong $\Theta_m \cap \mathcal{B}_1(0_p, 1)$. We also consider a positive radius r such that $(1 - \|\theta'\|_1 - 2rd_m)$ is positive. For any θ_1, θ_2 in $\mathcal{C}_m(\theta', r)$ the matrices $(I_{p^2} - C(\theta_1))$ and $(I_{p^2} - C(\theta_2))$ are diagonally dominant and their eigenvalues $\lambda_{1[i,j]}$ and $\lambda_{2[i,j]}$ are larger than $1 - \|\theta'\|_1 - 2rd_m$.

$$\begin{aligned} \mathcal{K}(\theta_1, \theta_2) &\leq \frac{9}{16(1 - \|\theta'\|_1 - 2rd_m)^2} \sum_{1 \leq i, j \leq p} (\lambda_{1[i,j]} - \lambda_{2[i,j]})^2 \\ &\leq \frac{9}{16(1 - \|\theta'\|_1 - 2rd_m)^2} \|C(\theta_1) - C(\theta_2)\|_F^2 \\ &\leq \frac{9d_m r^2 p^2}{8(1 - \|\theta'\|_1 - 2rd_m)^2}. \end{aligned} \quad (19)$$

We recall that $\|\cdot\|_F$ refers to the Frobenius norm in the space of matrices.

Let us state Birgé's version of Fano's lemma [2] and a combinatorial argument known under the name of Varshamov-Gilbert's lemma. These two lemma are taken from [4] and respectively correspond to Corollary 2.18 and Lemma 4.7.

Lemma 3.2. (Birgé's lemma) *Let (S, d) be some pseudo-metric space and $\{\mathbb{P}_s, s \in S\}$ be some statistical model. Let κ denote some absolute constant smaller than one. Then for any estimator \hat{s} and any finite subset T of S , setting $\delta = \min_{s, t \in T, s \neq t} d(s, t)$, provided that $\max_{s, t \in T} \mathcal{K}(\mathbb{P}_s, \mathbb{P}_t) \leq \kappa \log |T|$, the following lower bound holds for every $p \geq 1$,*

$$\sup_{s \in S} \mathbb{E}_s[d^p(s, \hat{s})] \geq 2^{-p} \delta^p (1 - \kappa) .$$

Lemma 3.3. (Varshamov-Gilbert's lemma) *Let $\{0, 1\}^d$ be equipped with Hamming distance d_H . There exists some subset Φ of $\{0, 1\}^d$ with the following properties*

$$d_H(\phi, \phi') > d/4 \text{ for every } (\phi, \phi') \in \Phi^2 \text{ with } \phi \neq \phi' \text{ and } \log |\Phi| \geq \frac{d}{8} .$$

Applying Lemma 3.2 with Hamming distance d_H and the set Φ introduced in Lemma 3.3 yields

$$\sup_{\theta \in \mathcal{C}_m(\theta', r)} \mathbb{E}_\theta \left[d_H(\hat{\theta}, \theta) \right] \geq \frac{d_m}{8} (1 - \kappa) , \quad (20)$$

provided that

$$\frac{9d_m r^2 p^2 n}{8(1 - \|\theta'\|_1 - 2rd_m)^2} \leq \frac{\kappa d_m}{8} . \quad (21)$$

Let us express (20) in terms of the Frobenius $\|\cdot\|_F$ norm.

$$\sup_{\theta \in \mathcal{C}_m(\theta', r)} \mathbb{E}_\theta \left[\|C(\hat{\theta}) - C(\theta)\|_F^2 \right] \geq \frac{d_m r^2 p^2}{4} (1 - \kappa) .$$

Since for every θ in the hypercube, $\sigma^{-2}(I_{p^2} - C(\theta))$ is diagonally dominant, its largest eigenvalue is smaller than $2\sigma^{-2}$. The loss function $l(\hat{\theta}, \theta)$ equals $\sigma^2/p^2 \text{tr}\{[C(\hat{\theta}) - C(\theta)](I - C(\theta))^{-1}[C(\hat{\theta}) - C(\theta)]\}$. It follows that

$$\sup_{\theta \in \mathcal{C}_m(\theta', r)} \mathbb{E}_\theta \left[l(\hat{\theta}, \theta) \right] \geq \sigma^2 \frac{d_m r^2}{8} (1 - \kappa) . \quad (22)$$

Condition (21) is equivalent to $r^2(1 - \|\theta'\|_1 - 2rd_m)^{-2} \leq \kappa/(9p^2n)$. If we assume that

$$r^2 \leq \frac{\kappa(1 - \|\theta'\|_1)^2}{18p^2n} , \quad (23)$$

then $1 - \|\theta'\|_1 - 2rd_m \geq (1 - \|\theta'\|_1) \left(1 - 2d_m \sqrt{\kappa/(18np^2)}\right)$. This last quantity is larger than $(1 - \|\theta'\|_1)/\sqrt{2}$ if d_m is smaller than $1.5(\sqrt{2}-1)\sqrt{np^2/\kappa}$. Gathering inequality (22) and condition (23), we get the lower bound

$$\begin{aligned} \inf_{\hat{\theta}} \sup_{\theta \in \text{Co}[\mathcal{C}_m(\theta', r)]} \mathbb{E}_{\theta} \left[l(\hat{\theta}, \theta) \right] &\geq \inf_{\hat{\theta}} \sup_{\theta \in \mathcal{C}_m \left[\theta', r \wedge (1 - \|\theta'\|_1) \sqrt{\frac{\kappa}{18p^2 n}} \right]} \mathbb{E}_{\theta} \left[l(\hat{\theta}, \theta) \right] \\ &\geq L \left(r^2 \wedge \frac{(1 - \|\theta'\|_1)^2}{np^2} \right) d_m \sigma^2. \end{aligned}$$

One handles models of dimension d_m between $1.5(\sqrt{2}-1)\sqrt{np^2/\kappa}$ and \sqrt{np} by changing the constant L in the last lower bound.

Let us turn to sets of isotropic GMRFs. The proof is similar to the non-isotropic case, except for a few arguments. Let m belongs to the collection \mathcal{M}_1 and let θ' be an element of $\Theta_m^{\text{iso}} \cap \mathcal{B}_1(0_p, 1)$. Let r be such that $1 - \|\theta'\|_1 - 8d_m^{\text{iso}}$ is positive. If θ_1 and θ_2 belong to the hypercube $\mathcal{C}_m^{\text{iso}}(\theta', r)$, then

$$\mathcal{K}(\theta_1, \theta_2) \leq \frac{9d_m r^2 p^2}{2(1 - \|\theta'\|_1 - 8rd_m^{\text{iso}})^2}.$$

Applying Lemma 3.2 and 3.3, it follows that

$$\inf_{\hat{\theta}} \sup_{\theta \in \mathcal{C}_m^{\text{iso}}(\theta', r)} \mathbb{E}_{\theta} \left[d_H(\hat{\theta}, \theta) \right] \geq \frac{d_m^{\text{iso}}}{8} (1 - \kappa),$$

provided that $4.5d_m r^2 p^2 n (1 - \|\theta'\|_1 - 8rd_m^{\text{iso}})^{-2} \leq \kappa d_m^{\text{iso}}/8$. As a consequence,

$$\inf_{\hat{\theta}} \sup_{\theta \in \mathcal{C}_m^{\text{iso}}(\theta', r)} \mathbb{E}_{\theta} \left[l(\hat{\theta}, \theta) \right] \geq \frac{d_m^{\text{iso}} r^2}{8} (1 - \kappa),$$

if $r^2 (1 - \|\theta'\|_1 - 8rd_m^{\text{iso}})^{-2} \leq \kappa(36p^2 n)^{-1}$. We conclude by arguing as in the isotropic case. \square

Proof of lemma 8.6 in [6]. Let m be a model in \mathcal{M}_1 , r be a positive number smaller than $1/(4d_m)$, and θ be an element of the convex hull of $\mathcal{C}_m(0_p, r)$. The covariance matrix of the vector X^v is $\Sigma = \sigma^2 [I - C(\theta)]^{-1}$. Since the field X is stationary, $\text{Var}_{\theta}(X_{[0,0]})$ equals any diagonal element of Σ . In particular, $\text{Var}_{\theta}(X_{[0,0]})$ corresponds to the mean of the eigenvalues of Σ . The matrix $(I - C(\theta))$ is block circulant. As in the proof of Lemma 20, we note λ the $p \times p$ matrix of the eigenvalues of $(I_{p^2} - C(\theta))$. By Lemma A.1 in [6],

$$\lambda_{[i,j]} = 1 + \sum_{(k,l) \in \Lambda} \theta_{[k,l]} \cos \left[2\pi \left(\frac{ik}{p} + \frac{jl}{p} \right) \right],$$

for any $1 \leq i, j \leq p$. Since θ belongs to the convex hull of $\mathcal{C}_m(0_p, r)$, $\theta_{[k,l]}$ is zero if $(k, l) \notin m$ and $|\theta_{[k,l]}| \leq r$ if $(k, l) \in m$. Thus $\sum_{(k,l) \in \Lambda} |\theta_{[k,l]}|$ is smaller than $1/2$. Applying Taylor-Lagrange inequality, we get

$$\frac{1}{1+x} \leq 1 - x + \frac{x^2}{(1-|x|)^3},$$

for any x between -1 and 1 . It follows that

$$\lambda_{[i,j]}^{-1} \leq 1 - \sum_{k,l \in \Lambda} \theta_{[k,l]} \cos \left[2\pi \left(\frac{ik}{p} + \frac{jl}{p} \right) \right] + 8 \left\{ \sum_{k,l \in \Lambda} \theta_{[k,l]} \cos \left[2\pi \left(\frac{ik}{p} + \frac{jl}{p} \right) \right] \right\}^2.$$

Summing this inequality for all $(i, j) \in \{1, \dots, p\}^2$, the first order term turns out to be $\text{tr}[C(\theta)]/p^2$ which is zero whereas the second term equals $8\text{tr}[C(\theta)^2]/p^2$. Since there are less than $2d_m$ non-zero terms on each line of the matrix $C(\theta)$, its Frobenius norm is smaller than $2d_m p^2 r^2$. Consequently, we obtain

$$\text{Var}_\theta(X_{[0,0]}) \leq \sigma^2 (1 + 16d_m r^2).$$

□

Proof of Lemma 8.7 in [6]. This property seems straightforward but the proof is a bit tedious. Let i be a positive integer smaller than $\text{Card}(\mathcal{M}_1)$. By definition of the radius r_m in Equation (10) in [6], the model m_{i+1} is the set of nodes in $\Lambda \setminus \{(0, 0)\}$ at a distance smaller or equal to $r_{m_{i+1}}$ from $(0, 0)$, whereas the model m_i only contains the points in $\Lambda \setminus \{(0, 0)\}$ at a distance strictly smaller than $r_{m_{i+1}}$ from the origin.

Let us first assume that $2r_{m_{i+1}} \leq p$. In such a case, the disc centered on $(0, 0)$ with radius $r_{m_{i+1}}$ does not overlap with itself on the torus Λ . To any node in the neighborhood m_{i+1} and to the node $(0, 0)$, we associate the square of size 1 centered on it. All these squares do not overlap and are included in the disc of radius $r_{m_{i+1}} + \sqrt{2}/2$. Hence, we get the upper bound $2d_{m_{i+1}} + 1 \leq \pi(r_{m_{i+1}} + \sqrt{2}/2)^2$. Similarly, the disc of radius $r_{m_{i+1}} - \sqrt{2}/2$ is included in the union of the squares associated to the nodes $m_i \cup \{0, 0\}$. It follows that $2d_{m_i} + 1$ is larger or equal to $\pi(r_{m_{i+1}} - \sqrt{2}/2)^2$. Gathering these two inequalities, we obtain

$$\frac{d_{m_{i+1}}}{d_{m_i}} \leq \frac{(r_{m_{i+1}} + \sqrt{2}/2)^2 - 1}{(r_{m_{i+1}} - \sqrt{2}/2)^2 - 1},$$

if $r_{m_{i+1}}$ is larger than $1 + \sqrt{2}/2$. If $r_{m_{i+1}}$ larger than 5, this upper bound is smaller than two. An exhaustive computation for models of small dimension allows to conclude.

If $2r_{m_{i+1}} \geq p$ and $2r_{m_i} < p$, then the preceding lower bound of d_{m_i} and the preceding upper bound of $d_{m_{i+1}}$ still hold. Finally, let us assume that $2r_{m_i} \geq p$. Arguing as previously, we conclude that $2d_{m_i} + 1 \geq \pi(p/2 - \sqrt{2}/2)^2$. The largest dimension of a model $m \in \mathcal{M}_1$ is $(p^2 - 1)/2$ if p is odd and $((p + 1)^2 - 3)/2$ if p is even. Thus, $d_{m_{i+1}} \leq [(p + 1)^2 - 3]/2$. Gathering these two bounds yields

$$\frac{d_{m_{i+1}}}{d_{m_i}} \leq 4 \frac{(p + 1)^2 - 3}{(p - \sqrt{2})^2},$$

which is smaller than 2 if p is larger than 10. Exhaustive computations for small p allow to conclude. \square

Proof of Proposition 6.7 in [6]. This result derives from the upper bound of the risk of $\tilde{\theta}_{\rho_1}$ stated in Theorem 3.1 and the minimax lower bound stated in Proposition 6.6 in [6].

Let $\mathcal{E}(a)$ be a pseudo-ellipsoid that satisfies Assumption (\mathbb{H}_a) and such that $a_1^2 \geq 1/(np^2)$. For any θ in $\mathcal{E}(a) \cap \mathcal{B}_1(0_p, 1) \cap \mathcal{U}(\rho_2)$, the penalty term satisfies $\text{pen}(m) = K\sigma^2\rho_1^2\rho_2 d_m / np^2$ is larger than $Kd_m\varphi_{\max}(\Sigma)/np^2$. Applying Theorem 3.1, we upper bound the risk θ_{ρ_1}

$$\mathbb{E}_\theta \left[l(\tilde{\theta}_{\rho_1}, \theta) \right] \leq L_1(K) \inf_{m \in \mathcal{M}_1} [l(\theta_{m, \rho_1}, \theta) + \text{pen}(m)] + L_2(K) \rho_2 \frac{\sigma^2}{np^2},$$

for any $\theta \in \mathcal{E}(a) \cap \mathcal{B}_1(0_p, 1) \cap \mathcal{U}(\rho_2)$. It follows that

$$\sup_{\theta \in \mathcal{E}(a) \cap \mathcal{B}_1(0_p, 1) \cap \mathcal{U}(\rho_2)} \mathbb{E}_\theta \left[l(\tilde{\theta}_{\rho_1}, \theta) \right] \leq L(K) \inf_{m \in \mathcal{M}_1, d_m > 0} \left[l(\theta_{m, \rho_1}, \theta) + \rho_1^2 \rho_2 \sigma^2 \frac{d_m}{np^2} \right].$$

Let i be a positive integer smaller or equal than $\text{Card}(\mathcal{M}_1)$. We know from Section 4.1 in [6] that the bias $l(\theta_{m_i}, \theta)$ of the model m_i equals $\text{Var}(X_{[0,0]} | X_{m_i}) - \sigma^2$. Since θ belongs to the set $\mathcal{E}(a) \cap \mathcal{B}_1(0_p, 1)$, the bias term is smaller or equal to a_{i+1}^2 with the convention $a_{\text{Card}(\mathcal{M}_1)+1}^2 = 0$. Hence, the previous upper bound becomes

$$\begin{aligned} \mathbb{E}_\theta \left[l(\tilde{\theta}_{\rho_1}, \theta) \right] &\leq L(K) \inf_{1 \leq i \leq \text{Card}(\mathcal{M}_1)} \left[a_{i+1}^2 + \rho_1^2 \rho_2 \sigma^2 \frac{d_{m_i}}{np^2} \right] \\ &\leq L(K, \rho_1, \rho_2) \inf_{1 \leq i \leq \text{Card}(\mathcal{M}_1)} \left[a_{i+1}^2 + \frac{\sigma^2 d_{m_i}}{np^2} \right]. \end{aligned} \quad (24)$$

Applying Proposition 6.6 in [6] to the set $\mathcal{E}(a) \cap \mathcal{B}_1(0_p, 1) \cap \mathcal{U}(2)$, we get

$$\begin{aligned} \inf_{\hat{\theta}} \sup_{\theta \in \mathcal{E}(a) \cap \mathcal{B}_1(0_p, 1) \cap \mathcal{U}(\rho_2)} \mathbb{E}_\theta \left[l(\hat{\theta}, \theta) \right] &\geq \inf_{\hat{\theta}} \sup_{\theta \in \mathcal{E}(a) \cap \mathcal{B}_1(0_p, 1) \cap \mathcal{U}(2)} \mathbb{E}_\theta \left[l(\hat{\theta}, \theta) \right] \\ &\geq L \sup_{1 \leq i \leq \text{Card}(\mathcal{M}_1)} \left(a_i^2 \wedge \sigma^2 \frac{d_{m_i}}{np^2} \right). \end{aligned}$$

Let us define i^* by

$$i^* := \sup \left\{ 1 \leq i \leq \text{Card}(\mathcal{M}_1), a_i^2 \geq \frac{\sigma^2 d_{m_i}}{np^2} \right\},$$

with the convention $\sup \emptyset = 0$. Since $a_1^2 \geq \sigma^2/np^2$, i^* is larger or equal to one. It follows that

$$\inf_{\widehat{\theta}} \sup_{\theta \in \mathcal{E}(a) \cap \mathcal{B}_1(0_p, \eta)} \mathbb{E}_{\theta} \left[l(\widehat{\theta}, \theta) \right] \geq L_2 \left(a_{i^*+1}^2 \vee \frac{\sigma^2 d_{m_{i^*}}}{np^2} \right).$$

Meanwhile, the upper bound (24) on the risk of $\widetilde{\theta}_{\rho_1}$ becomes

$$\mathbb{E}_{\theta} \left[l(\widetilde{\theta}_{\rho_1}, \theta) \right] \leq L(K, \rho_1, \rho_2) \left(a_{i^*+1}^2 + \frac{\sigma^2 d_{m_{i^*}}}{np^2} \right) \leq 2L(K, \rho_1, \rho_2) \left(a_{i^*+1}^2 \vee \frac{\sigma^2 d_{m_{i^*}}}{np^2} \right),$$

which allows to conclude. \square

4. Proof of the asymptotic risks bounds

Proof of Corollary 4.6 in [6]. For the sake of simplicity, we assume that for any node $(i, j) \in m$, the nodes (i, j) and $(-i, -j)$ are different in Λ . If this is not the case, we only have to slightly modify the proof in order to take account that $\|\Psi_{i,j}\|_F^2$ may equal one. The matrix V is the covariance of the vector of size d_m

$$(X_{i_1, j_1} + X_{-i_1, -j_1}, \dots, X_{i_{d_m}, j_{d_m}} + X_{-i_{d_m}, -j_{d_m}}). \quad (25)$$

Since the matrix Σ of X^v is positive, V is also positive. Moreover, its largest eigenvalue is larger than $2\varphi_{\max}(\Sigma)$.

Let us assume first the θ belongs to Θ_m^+ and that Assumption (\mathbb{H}_1) is fulfilled. By the first result of Proposition 4.4 in [6],

$$\lim_{n \rightarrow +\infty} np^2 \mathbb{E} \left[l(\widehat{\theta}_{m, \rho_1}, \theta) \right] = 2\sigma^4 \text{tr} [IL_m V^{-1}] \geq \frac{\sigma^4}{\varphi_{\max}(\Sigma)} \text{tr} [IL_m] = 2\sigma^4 \frac{d_m}{\varphi_{\max}(\Sigma)},$$

which corresponds to the first lower bound (30) in [6].

Let us turn to the second result. We now assume that θ satisfies Assumption (\mathbb{H}_2) . By the identity (28) of Proposition 4.4 in [6], we only have to lower bound the quantity $\text{tr} [VW^{-1}]$.

$$\text{tr} [V^{-1}W] \geq \varphi_{\max}(V)^{-1} \text{tr} [W] \geq \frac{1}{2\varphi_{\max}(\Sigma)} \text{tr} [W].$$

Since the matrix $\Sigma^{-1} = \sigma^{-2} [I_{p^2} - C(\theta)]$ is diagonally dominant, its smallest eigenvalue is larger than $\sigma^{-2}(1 - \|\theta\|_1)$. The matrix $(I_{p^2} - C(\theta_{m, \rho_1}))^2 (I_{p^2} - C(\theta))^{-2}$

is symmetric positive. It follows that W is also symmetric positive definite. Hence, we get

$$\begin{aligned} & \text{tr} [V^{-1}W] \\ & \geq \frac{\sigma^{-2}}{2} [1 - \|\theta\|_1] \sum_{k=1}^{d_m} \frac{\text{tr} [C(\Psi_{i_k, j_k})^2 [I_{p^2} - C(\theta_{m, \rho_1})]^2 [I_{p^2} - C(\theta)]^{-2}]}{p^2}. \end{aligned} \quad (26)$$

The largest eigenvalue of $(I_{p^2} - C(\theta))$ is smaller than 2 and the smallest eigenvalue of $(I_{p^2} - C(\theta_{m, \rho_1}))$ is larger than $1 - \|\theta_{m, \rho_1}\|_1$. By Lemma A.1 in [6], these two matrices are jointly diagonalizable and the smallest eigenvalue of

$$(I_{p^2} - C(\theta_{m, \rho_1}))^2 (I_{p^2} - C(\theta))^{-2}$$

is therefore larger than $(1 - \|\theta_{m, \rho_1}\|_1)^2/4$. Gathering this lower bound with (26) yields

$$\text{tr} [V^{-1}W] \geq \frac{d_m \sigma^{-2}}{2} [1 - \|\theta\|_1] [1 - \|\theta_{m, \rho_1}\|_1]^2.$$

Lemma 4.1 in [6] states that $\|\theta_{m, \rho_1}\|_1 \leq \|\theta\|_1$. Combining these two lower bounds enables to conclude. \square

Proof of Example 4.8 in [6].

Lemma 4.1. *For any θ in the space $\Theta_{m_1}^{+, \text{iso}}$, the asymptotic variance term of $\hat{\theta}_{m_1, \rho_1}^{\text{iso}}$ equals*

$$\lim_{n \rightarrow +\infty} np^2 \mathbb{E}_\theta \left[l \left(\hat{\theta}_{m_1, \rho_1}^{\text{iso}}, \theta \right) \right] = 2\sigma^4 \frac{\text{tr} (H^2)}{\text{tr} (H^2 \Sigma)}.$$

If θ belongs to $\Theta^{+, \text{iso}}$ and also satisfies (\mathbb{H}_2) , then

$$\lim_{n \rightarrow +\infty} np^2 \mathbb{E}_\theta \left[l \left(\hat{\theta}_{m_1, \rho_1}^{\text{iso}}, \theta_{m_1, \rho_1}^{\text{iso}} \right) \right] = 2 \frac{\text{tr} \left\{ [(I - \theta_{m_1, \rho_1}^{\text{iso}} [1, 0] H) H \Sigma]^2 \right\}}{\text{tr} (H^2 \Sigma)}, \quad (27)$$

where the $p^2 \times p^2$ matrix H is defined as $H := C(\Psi_{1,0}^{\text{iso}})$.

Proof of Lemma 4.1. Apply Proposition 4.4 in [6] noting that $V = \text{tr}[H \Sigma H]/p^2$ and

$$W = \frac{\text{tr} \left\{ [(I - \theta_{m_1, \rho_1}^{\text{iso}} [1, 0] H) H \Sigma]^2 \right\}}{\sigma^4 p^2}.$$

To prove the second result, we observe that $\Theta_{m_1}^{+, \text{iso}}$ equals $\Theta_{m_1, 2}^{+, \text{iso}}$. It is stated for instance in Table 2 in [6]. \square

Since the matrix θ belongs to $\Theta_{m_1}^{+, \text{iso}}$, we may apply the second result of Lemma 4.1. Straightforward computations lead to $\text{tr}(H^2) = \|C(\Psi_{1,0}^{\text{iso}})\|_F^2 = 4p^2$ and

$$\text{tr}(H^2\Sigma) = 4p^2 [\text{Var}(X_{[0,0]}) + 2\text{cov}_\theta(X_{[0,0]}, X_{[1,1]}) + \text{cov}_\theta(X_{[0,0]}, X_{[2,0]})] .$$

Since the field X is an isotropic GMRF with four nearest neighbors,

$$X_{[0,0]} = \theta_{[1,0]} (X_{[1,0]} + X_{[-1,0]} + X_{[0,1]} + X_{[0,-1]}) + \epsilon_{[0,0]} ,$$

where $\epsilon_{[0,0]}$ is independent from every variable $X_{[i,j]}$ with $(i,j) \neq 0$. Multiplying this identity by $X_{[1,0]}$ and taking the expectation yields

$$\text{cov}_\theta(X_{[0,0]}, X_{[1,0]}) = \theta_{[1,0]} [\text{Var}(X_{[0,0]}) + 2\text{cov}_\theta(X_{[0,0]}, X_{[1,1]}) + \text{cov}_\theta(X_{[0,0]}, X_{[2,0]})] .$$

Hence, we obtain $\text{tr}(H^2\Sigma) = 4\text{cov}_\theta(X_{[0,0]}, X_{[1,0]})/\theta_{[1,0]}$ and

$$\frac{\text{tr}(H^2)}{\text{tr}(H^2\Sigma)} = \frac{\theta_{[1,0]}}{\text{cov}_\theta(X_{[0,0]}, X_{[1,0]})} ,$$

which concludes the first part of the proof.

This second part is based on the spectral representation of the field X and follows arguments which come back to Moran [5]. We shall compute the limit of $\text{cov}_\theta(X_{[0,0]}, X_{[1,0]})$ when the size of Λ goes to infinity. As the field X is stationary on Λ , we may diagonalize its covariance matrix Σ applying Lemma A.1 in [6]. We note D_Σ the corresponding diagonal matrix defined by

$$D_{\Sigma[(i-1)p+j, (i-1)p+j]} = \sum_{k=1}^p \sum_{l=1}^p \text{cov}_\theta(X_{[0,0]}, X_{[k,l]}) \cos \left[2\pi \left(\frac{ki}{p} + \frac{lj}{p} \right) \right] ,$$

for any $1 \leq i, j \leq p$. Straightforwardly, we express $\text{cov}_\theta(X_{[0,0]}, X_{[1,0]})$ as a linear combination of the eigenvalues

$$\text{cov}_\theta(X_{[0,0]}, X_{[1,0]}) = \frac{1}{p^2} \sum_{i=1}^p \sum_{j=1}^p \cos \left(2\pi \frac{i}{p} \right) D_{\Sigma[(i-1)p+j, (i-1)p+j]} .$$

Applying Lemma A.1 in [6] to the matrix Σ^{-1} and noting that $\theta \in \Theta^{\text{iso}, +}$ allows to get another expression of the eigenvalues of Σ

$$D_{\Sigma[(i-1)p+j, (i-1)p+j]} = \frac{\sigma^2}{1 - 2\theta_{[1,0]} \left[\cos \left(\frac{2\pi i}{p} \right) + \cos \left(\frac{2\pi j}{p} \right) \right]} .$$

We then combine these expression. By symmetry between i and j we get

$$\text{cov}_\theta(X_{[0,0]}, X_{[1,0]}) = \frac{\sigma^2}{2p^2} \sum_{i=1}^p \sum_{j=1}^p \frac{\cos \left(2\pi \frac{i}{p} \right) + \cos \left(2\pi \frac{j}{p} \right)}{1 - 2\theta_{[1,0]} \left[\cos \left(2\pi \frac{i}{p} \right) + \cos \left(2\pi \frac{j}{p} \right) \right]} .$$

If we let p go to infinity, this sum converges to the following integral

$$\begin{aligned} & \lim_{p \rightarrow +\infty} \text{cov}_\theta (X_{[0,0]}, X_{[1,0]}) \\ &= \frac{\sigma^2}{2} \int_0^1 \int_0^1 \frac{\cos(2\pi x) + \cos(2\pi y)}{1 - 2\theta_{[1,0]} (\cos(2\pi x) + \cos(2\pi y))} dx dy \\ &= \frac{\sigma^2}{2\theta_{[1,0]}} \left[-1 + \frac{1}{4\pi^2} \int_0^{2\pi} \int_0^{2\pi} \frac{1}{1 - 2\theta_{[1,0]} [\cos(x) + \cos(y)]} dx dy \right]. \end{aligned}$$

This last elliptic integral is asymptotically equivalent to $\log 16[4(1 - 4\theta_{[1,0]})]^{-1}$ when $\theta_{[1,0]} \rightarrow 1/4$ as observed for instance by Moran [5]. We conclude by substituting this limit in expression (33) in [6]. \square

Proof of Example 4.9 in [6]. First, we compute $[\theta^{(p)}]_{m_1}^{\text{iso}}[1,0]$. By Lemma 4.1 in [6], it minimizes the function $\gamma(\cdot)$ defined in (19) in [6] over the whole space $\Theta_{m_1}^{\text{iso}}$. We therefore obtain

$$[\theta^{(p)}]_{m_1}^{\text{iso}}[1,0] = \frac{\text{tr}[\Sigma H]}{\text{tr}[\Sigma H^2]}.$$

Once again, we apply Lemma A.1 in [6] to simultaneously diagonalize the matrices H and Σ^{-1} . As previously, we note D_Σ the corresponding diagonal matrix of Σ .

$$\begin{aligned} D_{\Sigma[(i-1)p+j, (i-1)p+j]} &= \frac{\sigma^2}{1 - 2\alpha \left[\cos\left(2\pi\left(\frac{pi}{4p} + \frac{pj}{4p}\right)\right) + \cos\left(2\pi\left(\frac{-pi}{4p} + \frac{pj}{4p}\right)\right) \right]} \\ &= \frac{\sigma^2}{1 - 4\alpha \cos\left(\pi\frac{i}{2}\right) \cos\left(\pi\frac{j}{2}\right)}. \end{aligned}$$

Analogously, we compute the diagonal matrix $D(\Psi_{1,0}^{\text{iso}})$

$$D(\Psi_{1,0}^{\text{iso}})[(i-1)p+j, (i-1)p+j] = 2 \left[\cos\left(2\pi\frac{i}{p}\right) + \cos\left(2\pi\frac{j}{p}\right) \right].$$

Combining these two last expressions, we obtain

$$\text{tr}(H\Sigma) = \sum_{i=1}^p \sum_{j=1}^p \sigma^2 \frac{2 \left[\cos\left(2\pi\frac{i}{p}\right) + \cos\left(2\pi\frac{j}{p}\right) \right]}{1 - 4\alpha \cos\left(\pi\frac{i}{2}\right) \cos\left(\pi\frac{j}{2}\right)}.$$

Let us split this sum in 16 parts depending on the congruence of i and j modulo 4. As each if of these 16 sums is shown to be zero, we conclude that $\text{tr}(H\Sigma) = [\theta^{(p)}]_{m_1}^{\text{iso}}[1,0] = 0$. By Lemma 4.1, the asymptotic risk of $\widehat{\theta^{(p)}}_{m_1}^{\text{iso}, \rho_1}$ therefore equals

$$\lim_{n \rightarrow +\infty} np^2 \mathbb{E}_{\theta^{(p)}} \left[l \left(\widehat{\theta^{(p)}}_{m_1}^{\text{iso}, \rho_1}, [\theta^{(p)}]_{m_1}^{\text{iso}} \right) \right] = \frac{\text{tr}(H^4 \Sigma^2)}{\text{tr}(H^2 \Sigma)}.$$

First, we lower bound the numerator

$$\text{tr}(H^4 \Sigma^2) = \sigma^4 \sum_{i=1}^p \sum_{j=1}^p \frac{\left\{ 2 \left[\cos \left(2\pi \frac{i}{p} \right) + \cos \left(2\pi \frac{j}{p} \right) \right] \right\}^4}{\left\{ 1 - 4\alpha \cos \left(\pi \frac{i}{2} \right) \cos \left(\pi \frac{j}{2} \right) \right\}^2}.$$

As each term of this sum is non-negative, we may only consider the coefficients i and j which are congruent to 0 modulo 4.

$$\text{tr}(H^4 \Sigma^2) \geq \sigma^4 \sum_{i=0}^{p/4-1} \sum_{j=0}^{p/4-1} \frac{16 \left[\cos \left(2\pi \frac{i}{p/4} \right) + \cos \left(2\pi \frac{j}{p/4} \right) \right]^4}{(1 - 4\alpha)^2}.$$

If we let go p to infinity, we get the lower bound

$$\lim_{p \rightarrow +\infty} \frac{\text{tr}(H^4 \Sigma^2)}{p^2} \geq \frac{\sigma^4}{(1 - 4\alpha)^2} \int_0^1 \int_0^1 [\cos(2\pi x) + \cos(2\pi y)]^4 dx dy.$$

Similarly, we upper bound $\text{tr}(H^2 \Sigma)$ and let p go to infinity

$$\lim_{p \rightarrow +\infty} \frac{\text{tr}(H^2 \Sigma)}{p^2} \leq \frac{4\sigma^2}{1 - 4\alpha} \int_0^1 \int_0^1 [\cos(2\pi x) + \cos(2\pi y)]^2 dx dy.$$

Combining these two bounds allows to conclude

$$\lim_{p \rightarrow +\infty} \lim_{n \rightarrow +\infty} np^2 R_{\theta(p)} \left(\widehat{\theta^{(p)}}_{m_1}^{\text{iso}, \rho_1}, [\theta^{(p)}]_{m_1}^{\text{iso}} \right) \geq \frac{L\sigma^2}{1 - 4\alpha}.$$

□

5. Miscellaneous

Proof of Lemma 1.1 in [6]. Let θ be a $p \times p$ matrix that satisfies condition (3) in [6]. For any $1 \leq i_1, i_2 \leq p$, we define the $p \times p$ submatrix C_{i_1, i_2} as

$$C_{i_1, i_2}[j_1, j_2] := C(\theta)[(i_1-1)p+j_1, (i_2-1)p+j_2],$$

for any $1 \leq j_1, j_2 \leq p$. For the sake of simplicity, the subscripts (i_1, i_2) are taken modulo p . By definition of $C(\theta)$, it holds that $C_{i_1, i_2} = C_{0, i_2-i_1}$ for any $1 \leq i_1, i_2 \leq p$. Besides, the matrices $C_{0, i}$ are circulant for any $1 \leq i \leq p$. In short, the matrix $C(\theta)$ is of the form

$$C(\theta) = \begin{pmatrix} C_{0,1} & C_{0,2} & \cdots & C_{0,p} \\ \vdots & \vdots & \vdots & \vdots \\ C_{0,p} & C_{0,1} & \cdots & C_{0,p-1} \end{pmatrix},$$

where the matrices $C_{0, i}$ are circulant. Let (i_1, i_2, j_1, j_2) be in $\{1, \dots, p\}^4$. By definition,

$$C(\theta)[(i_1-1)p+j_1, (i_2-1)p+j_2] = \theta_{[i_2-i_1, j_2-j_1]}.$$

Since the matrix θ satisfies condition (3) in [6], $\theta_{[i_2-i_1, j_2-j_1]} = \theta_{[i_1-i_2, j_1-j_2]}$. As a consequence,

$$C(\theta)_{[(i_1-1)p+j_1, (i_2-1)p+j_2]} = C(\theta)_{[(i_2-1)p+j_2, (i_1-1)p+j_1]} \text{ and } C(\theta) \text{ is symmetric.}$$

Conversely, let B be a $p^2 \times p^2$ symmetric block circulant matrix. Let us define the matrix θ of size p by

$$\theta_{[i,j]} := B_{[1, (i-1)p+j]} ,$$

for any $1 \leq i, j \leq p$. Since the matrix B is block circulant, it follows that $C(\theta) = B$. By definition, $\theta_{[i,j]} = C(\theta)_{[1, (i-1)p+j]}$ and $\theta_{[-i, -j]} = C(\theta)_{[(i-1)p+j, 1]}$ for any integers $1 \leq i, j \leq p$. Since the matrix B is symmetric, we conclude that $\theta_{[i,j]} = \theta_{[-i, -j]}$. \square

Proof of Lemma 2.2 in [6]. For any $\theta' \in \Theta^+$, $\gamma_{n,p}(\theta')$ is defined as

$$\gamma_{n,p}(\theta') = \frac{1}{p^2} \text{tr} \left[(I_{p^2} - C(\theta')) \overline{\mathbf{X}^v \mathbf{X}^{v*}} (I_{p^2} - C(\theta')) \right] .$$

Applying Lemma A.1 in [6], there exists an orthogonal matrix P that simultaneously diagonalizes Σ and any matrix $C(\theta')$. Let us define $\mathbf{Y}^i := \sqrt{\Sigma}^{-1} \mathbf{X}_i$ and $D_\Sigma := P \Sigma P^*$. Gathering these new notations yields

$$\gamma_{n,p}(\theta') = \frac{1}{p^2} \text{tr} \left[(I_{p^2} - D(\theta')) D_\Sigma \overline{\mathbf{Y} \mathbf{Y}^*} (I_{p^2} - D(\theta')) \right] ,$$

where the vectors \mathbf{Y}^i are independent standard Gaussian random vectors. Except $\overline{\mathbf{Y} \mathbf{Y}^*}$, every matrix involved in this last expression is diagonal. Besides, the diagonal matrix D_Σ is positive since Σ is non-singular. Thus,

$$\text{tr} \left[(I_{p^2} - D(\theta')) D_\Sigma \overline{\mathbf{Y} \mathbf{Y}^*} (I_{p^2} - D(\theta')) \right]$$

is almost surely a positive quadratic form on the vector space generated by I_{p^2} and $D(\Theta^+)$. Since the function $D(\cdot)$ is injective and linear on Θ^+ , it follows that $\gamma_{n,p}(\cdot)$ is almost surely strictly convex on Θ^+ . \square

Proof of Lemma 4.1 and Corollary 4.2 in [6]. The proof only uses the stationarity of the field X on Λ and the l_1 norm of θ . However, the computations are a bit cumbersome. Let θ be an element of Θ^+ . By standard Gaussian properties, the expectation of $X_{[0,0]}$ given the remaining covariates is

$$\mathbb{E}_\theta \left(X_{[0,0]} | X_{-\{0,0\}} \right) = \sum_{(i,j) \in \Lambda \setminus (0,0)} \theta_{[i,j]} X_{[i,j]} .$$

By assumption (\mathbb{H}_2) , the l_1 norm of θ is smaller than one. We shall prove by backward induction that for any subset A of $\Lambda \setminus \{(0,0)\}$ the matrix θ^A uniquely defined by

$$\mathbb{E}_\theta (X_{[0,0]}|X_A) = \sum_{(i,j) \in A} \theta^A_{[i,j]} X_{[i,j]} \text{ and } \theta^A_{[i,j]} = 0 \text{ for any } (i,j) \notin A$$

satisfies $\|\theta^A\|_1 \leq \|\theta\|_1$. The property is clearly true if $A = \Lambda \setminus \{(0,0)\}$. Suppose we have proved it for any set of cardinality q larger than one. Let A be a subset of $\Lambda \setminus \{(0,0)\}$ of cardinality $q-1$ and (i,j) be an element of $\Lambda \setminus (A \cup \{(0,0)\})$. Let us derive the expectation of $X_{[0,0]}$ conditionally to X_A from the expectation of $X_{[0,0]}$ conditionally to $X_{A \cup \{(i,j)\}}$.

$$\begin{aligned} \mathbb{E}_\theta (X_{[0,0]}|X_A) &= \mathbb{E}_\theta [\mathbb{E}(X_{[0,0]}|X_A)|X_{A \cup \{(i,j)\}}] \\ &= \sum_{(k,l) \in A} \theta^{A \cup \{(i,j)\}}_{[k,l]} X_{[k,l]} + \theta^{A \cup \{(i,j)\}}_{[i,j]} \mathbb{E}_\theta [X_{[i,j]}|X_A] \end{aligned} \quad (28)$$

Let us take the conditional expectation of $X_{[i,j]}$ with respect to $X_{A \cup \{(0,0)\}}$. Since the field X is stationary on Λ and by the induction hypothesis, the unique matrix $\theta^{A \cup \{(0,0)\}}_{(i,j)}$ defined by

$$\mathbb{E}_\theta (X_{[i,j]}|X_{A \cup \{(0,0)\}}) = \sum_{(k,l) \in A \cup \{(0,0)\}} \theta^{A \cup \{(0,0)\}}_{(i,j)}_{[k,l]} X_{[k,l]}$$

and $\theta^{A \cup \{(0,0)\}}_{(i,j)}_{[k,l]} = 0$ for any $(k,l) \notin A \cup \{(0,0)\}$ satisfies $\|\theta^{A \cup \{(0,0)\}}_{(i,j)}\|_1 \leq \|\theta\|_1$. Taking the expectation conditionally to X_A of this previous expression leads to

$$\mathbb{E}_\theta (X_{[i,j]}|X_A) = \sum_{(k,l) \in A} \theta^{A \cup \{(0,0)\}}_{(i,j)}_{[k,l]} X_{[k,l]} + \theta^{A \cup \{(0,0)\}}_{(i,j)}_{[0,0]} \mathbb{E}_\theta (X_{[0,0]}|X_A) . \quad (29)$$

Gathering identities (28) and (29) yields

$$\mathbb{E}_\theta (X_{[0,0]}|X_A) = \sum_{(k,l) \in A} \frac{\theta^{A \cup \{(i,j)\}}_{[k,l]} + \theta^{A \cup \{(i,j)\}}_{[i,j]} \theta^{A \cup \{(0,0)\}}_{(i,j)}_{[k,l]}}{1 - \theta^{A \cup \{(i,j)\}}_{[i,j]} \theta^{A \cup \{(0,0)\}}_{(i,j)}_{[0,0]}} X_{[k,l]} ,$$

since $|\theta^{A \cup \{(i,j)\}}_{[i,j]} \theta^{A \cup \{(0,0)\}}_{(i,j)}_{[0,0]}| < 1$. Then, we upper bound the l_1 norm of θ^A

using that $\|\theta^{A \cup \{(i,j)\}}\|_1$ and $\|\theta_{(i,j)}^{A \cup \{(0,0)\}}\|_1$ are smaller or equal to $\|\theta\|_1$.

$$\begin{aligned}
& \|\theta^A\|_1 \\
& \leq \frac{\sum_{(k,l) \in A} |\theta^{A \cup \{j+1\}}_{[k,l]}| + \sum_{(k,l) \in A} |\theta^{A \cup \{(i,j)\}}_{[i,j]} \theta_{(i,j)}^{A \cup \{(0,0)\}}_{[k,l]}|}{1 - |\theta^{A \cup \{(i,j)\}}_{[i,j]} \theta_{(i,j)}^{A \cup \{(0,0)\}}_{[0,0]}|} \\
& \leq \frac{\|\theta\|_1 + |\theta^{A \cup \{(i,j)\}}_{[i,j]}| \left(\sum_{(k,l) \in A \cup \{(0,0)\}} |\theta_{(i,j)}^{A \cup \{(0,0)\}}_{[k,l]}| - 1 - |\theta_{(i,j)}^{A \cup \{(0,0)\}}_{[0,0]}| \right)}{1 - |\theta^{A \cup \{(i,j)\}}_{[i,j]} \theta_{(i,j)}^{A \cup \{(0,0)\}}_{[0,0]}|} \\
& \leq \frac{\|\theta\|_1 (1 + |\theta^{A \cup \{(i,j)\}}_{[i,j]}|) - |\theta^{A \cup \{(i,j)\}}_{[i,j]}| (1 + |\theta_{(i,j)}^{A \cup \{(0,0)\}}_{[0,0]}|)}{1 - |\theta^{A \cup \{(i,j)\}}_{[i,j]} \theta_{(i,j)}^{A \cup \{(0,0)\}}_{[0,0]}|} \\
& \leq \|\theta\|_1 + \frac{|\theta^{A \cup \{(i,j)\}}_{[i,j]}| (\|\theta\|_1 - 1) (1 + |\theta_{(i,j)}^{A \cup \{(0,0)\}}_{[0,0]}|)}{1 - |\theta^{A \cup \{(i,j)\}}_{[i,j]} \theta_{(i,j)}^{A \cup \{(0,0)\}}_{[0,0]}|}.
\end{aligned}$$

Since $\|\theta\|_1$ is smaller than one, it follows that $\|\theta^A\|_1 \leq \|\theta\|_1$.

Let m be a model in the collection \mathcal{M}_1 . Since m stands for a set of neighbors of $(0,0)$, we may define θ^m as above. It follows that $\|\theta^m\|_1 \leq \|\theta\|_1$. Since the field X is stationary on the torus, X follows the same distribution as the field X^s defined by $X^s_{[i,j]} = X_{[-i,-j]}$. By uniqueness of θ^m , we obtain that $\theta^m_{[i,j]} = \theta^m_{[-i,-j]}$. Thus, θ^m belongs to the space Θ_m . Moreover, θ^m minimizes the function $\gamma(\cdot)$ on Θ_m . Since the l_1 norm of θ^m is smaller than one, θ^m belongs to $\Theta_{m,2}^+$. The matrices θ^m and θ_{m,ρ_1} are therefore equal, which concludes the proof in the non-isotropic case.

Let us now turn to the isotropic case. Let θ belong to $\Theta^{\text{iso},+}$ and let m be a model in \mathcal{M}_1 . As previously, the matrix θ^m satisfies $\|\theta^m\|_1 \leq \|\theta\|_1$. Since the distribution of X is invariant under the action of the group G , θ^m belongs to Θ_m^{iso} . Since $\|\theta^m\|_1 \leq \|\theta\|_1$, θ^m lies in $\Theta_{m,2}^{+, \text{iso}}$. It follows that $\theta^m = \theta_{m,\rho_1}^{\text{iso}}$. \square

Proof of Corollary 4.3 in [6]. Let θ be a matrix in Θ^+ such that (\mathbb{H}_2) holds and let m be a model in \mathcal{M}_1 . We decompose $\gamma(\widehat{\theta}_{m,\rho_1})$ using the conditional

expectation of $X_{[0,0]}$ given X_m .

$$\begin{aligned}
\gamma(\widehat{\theta}_{m,\rho_1}) &= \mathbb{E}_\theta \left[X_{[0,0]} - \sum_{(i,j) \in m} \widehat{\theta}_{m,\rho_1}[i,j] X_{[i,j]} \right]^2 \\
&= \mathbb{E}_\theta \left[X_{[0,0]} - \mathbb{E}_\theta (X_{[0,0]} | X_m) \right]^2 \\
&+ \mathbb{E}_\theta \left[\mathbb{E}_\theta (X_{[0,0]} | X_m) - \sum_{(i,j) \in m} \widehat{\theta}_{m,\rho_1}[i,j] X_{[i,j]} \right]^2.
\end{aligned}$$

By Corollary (11) in [6], we know that

$$\mathbb{E}_\theta (X_{[0,0]} | X_m) = \sum_{(i,j) \in m} \theta_{m,\rho_1}[i,j] X_{[i,j]}.$$

Combining these two last identities yields

$$\gamma(\widehat{\theta}_{m,\rho_1}) = \gamma(\theta_{m,\rho_1}) + \mathbb{E}_\theta \left[\sum_{(i,j) \in \Lambda \setminus \{(0,0)\}} (\theta_{m,\rho_1} - \widehat{\theta}_{m,\rho_1})[i,j] X_{[i,j]} \right]^2.$$

Subtracting $\gamma(\theta)$, we obtain the first result. The proof is analogous in the isotropic case. \square

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